



On the optimality of
satellite-based CO₂
atmospheric
inversions

F. Chevallier

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On the statistical optimality of CO₂ atmospheric inversions assimilating CO₂ column retrievals

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Abstract

The extending archive of the Greenhouse Gases Observing SATellite (GOSAT) measurements (now covering about six years) allows increasingly robust statistics to be computed, that document the performance of the corresponding retrievals of the column-average dry air-mole fraction of CO₂ (XCO_2). Here, we compare a model simulation constrained by surface air-sample measurements with one of the GOSAT retrieval products (NASA's ACOS). The retrieval-minus-model differences result from various error sources, both in the retrievals and in the simulation: we discuss the plausibility of the origin of the major patterns. We find systematic retrieval errors over the dark surfaces of high-latitude lands and over African savannahs. More importantly, we also find a systematic over-fit of the GOSAT radiances by the retrievals over land for the high-gain detector mode, which is the usual observation mode. The over-fit is partially compensated by the retrieval bias-correction. These issues are likely common to other retrieval products and may explain some of the surprising and inconsistent CO₂ atmospheric inversion results obtained with the existing GOSAT retrieval products. We suggest that reducing the observation weight in the retrieval schemes (for instance so that retrieval increments to the retrieval prior values are halved for the studied retrieval product) would significantly improve the retrieval quality and reduce the need for (or at least reduce the complexity of) ad-hoc retrieval bias correction. More generally, we demonstrate that atmospheric inversions cannot be rigorously optimal when assimilating XCO_2 retrievals, even with averaging kernels.

1 Introduction

CO₂ surface fluxes at the Earth's surface can be inferred from accurate surface measurements of CO₂ concentrations, but the sparseness of the current global network still leaves the flux horizontal and temporal gradients, and even their latitudinal distribution, very uncertain (Peylin et al., 2013). This limitation has provided a major in-

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(MOE) in Japan. This spacecraft is operated in a sun-synchronous polar orbit that crosses the Equator at about 13:00 LT during daytime and that repeats every 3 days. As described by O'Dell et al. (2012) and Osterman et al. (2013), the ACOS algorithm retrieves XCO_2 from a selection of GOSAT measurements of reflected sunlight made in the same spectral bands than OCO-2. Over land, such measurements are made by pointing the instrument to the Earth on both sides of the satellite track. Given the low reflectivity of water surfaces, ocean measurements are only possible when the instrument is pointed to the sun-glint spot, which is only done within 40° from the Equator in the summer hemisphere. GOSAT also carries a cloud and aerosol imager that can help filtering difficult scenes out, but unlike other GOSAT retrieval algorithms, ACOS does not use it since OCO-2 does not contain a similar instrument.

Following Boesch et al. (2006) and Connor et al. (2008), the ACOS algorithm relies on optimal estimation (i.e. Bayesian methods) to retrieve the vertical profile of the CO₂ dry air mole fraction together with variables interfering in the measurements: the surface pressure and the surface albedo, some variables describing temperature, water vapour, clouds and aerosols in the atmosphere, and channel offsets for the instrument. The retrieved XCO_2 is simply obtained by integrating the retrieved CO₂ profile. In this Bayesian formulation of the retrieval, prior information about CO₂ is given an artificially small weight in order to maximize the observation contribution to the result: for instance, the standard deviation of the uncertainty assigned to the prior XCO_2 is larger than 10 ppm (O'Dell et al., 2012), i.e. larger than typical variations of XCO_2 at the continental scale (e.g., Keppel-Aleks et al., 2011). We will discuss the impact of this choice later and for simplicity, we will call XCO_2^b and XCO_2^a the prior (*background*) and the retrieved (*analysed*) XCO_2 , respectively. XCO_2^a can be compared with model simulations, as will be done here, or with other measurements via the associated CO₂ averaging kernel profiles and prior profiles (e.g., Connor et al., 1994). For nadir viewing, XCO_2^a is representative of a volume that has a circular footprint at the Earth's surface of diameter about 10 km.

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Previous comparisons between XCO_2^a and model simulations or reference ground-based XCO_2 measurements from Total Carbon Column Observing Network (TCCON) highlighted some systematic dependency of the error of XCO_2^a as a function of a series of internal variables of the algorithm (Wunch et al., 2011b). This feature reveals some limitations of the algorithm but also allows correcting them empirically, for instance before they are assimilated in atmospheric inversion systems (Crisp et al., 2012). We will call $XCO_2^{a,c}$ the bias-corrected retrievals.

2.2 MACC CO₂ inversion

Since year 2011, the MACC pre-operational service (www.copernicus-atmosphere.eu) has been delivering a CO₂ inversion product with biannual updates. Release 13r1 primarily describes the CO₂ surface fluxes over more than three decades, from 1979 to 2013, at resolution $3.75^\circ \times 1.9^\circ$ (longitude–latitude) and 3 hourly, based on 131 CO₂ dry air mole fraction station records from three large databases:

- the NOAA Earth System Research Laboratory archive (NOAA CCGG, <http://www.esrl.noaa.gov/gmd/ccgg/index.html>),
- the World Data Centre for Greenhouse Gases archive (WDCGG, <http://ds.data.jma.go.jp/gmd/wdcgg/>),
- the Réseau Atmosphérique de Mesure des Composés à Effet de Serre database (RAMCES, <http://www.lsce.ipsl.fr/>).

The three databases include both in situ measurements made by automated quasi-continuous analysers and irregular air samples collected in flasks and later analyzed in central facilities. The detailed list of sites is provided in Tables 1 and 2.

The MACC Bayesian inversion method is formulated in a variational way in order to estimate the CO₂ surface fluxes at the above-described relatively high resolution over the globe (Chevallier et al., 2005, 2010). For v13r1, the system used a single 35 year inversion window, therefore enforcing physical and statistical consistency in

written:

$$\mathbf{x}^a = \mathbf{x}^b + \mathbf{K}(\mathbf{y} - \mathbf{H}\mathbf{x}^b) \quad (1)$$

\mathbf{H} a linearized observation operator that links variables \mathbf{x} and \mathbf{y} (i.e. essentially a radiative transfer model or a transport model). \mathbf{K} is the following “Kalman gain” matrix:

$$\mathbf{K} = \mathbf{B}\mathbf{H}^T(\mathbf{H}\mathbf{B}\mathbf{H}^T + \mathbf{R})^{-1} \quad (2)$$

\mathbf{B} and \mathbf{R} are the error covariance matrices of \mathbf{x}^b and \mathbf{y} , respectively.

The error covariance matrix of \mathbf{x}^a is obtained by:

$$\mathbf{A} = (\mathbf{I} - \mathbf{K}\mathbf{H})\mathbf{B} \quad (3)$$

with \mathbf{I} the identity matrix with appropriate dimension.

For simplicity, Eq. (1) does not make other variables that are simultaneously inferred appear, like clouds, aerosols or surface variables for the retrievals, or the 3-D state of CO_2 at the start of the inversion window for the inversion.

The current processing chains that go from radiances to surface fluxes are two-step processes. We now distinguish the retrieval process and the inversion process by putting breves \checkmark on all symbols related to the former and hats $\hat{}$ on all symbols related to the latter. In a first step, the CO_2 profiles and their uncertainty $\checkmark\mathbf{x}$, $\checkmark\mathbf{A}$ are retrieved for each sounding separately. The resulting ensemble forms the observations to be simultaneously assimilated $\hat{\mathbf{y}}$, $\hat{\mathbf{R}}$ for the second step. The presence of prior information \mathbf{x}^b in both steps complicates the transition between the two. Following Connor et al. (1994), we can technically eliminate the influence of $\checkmark\mathbf{x}^b$ (but not of its uncertainty) by the following adaptation of Eq. (1) in the second step: we subtract the retrieval prior $\checkmark\mathbf{x}^b$ from each CO_2 profile simulated by the transport model at the sounding location $\hat{\mathbf{H}}\hat{\mathbf{x}}^b$, we multiply the result by the retrieval averaging kernel matrix $\checkmark\checkmark\mathbf{H}$ and finally add $\checkmark\mathbf{x}^b$. The retrieval error covariance matrix should consistently be diminished (e.g., Connor et al., 2008, paragraph 37). We call $\hat{\mathbf{H}}'$ the convolution of the transport model operator with

the individual retrieval averaging kernels and $\hat{\mathbf{R}}'$ the adjusted retrieval error covariance matrix. By applying Eq. (1) twice and after accounting for the above adaptation, the processing chain can be written in a concise form:

$$\hat{\mathbf{x}}^a = \hat{\mathbf{x}}^b + \hat{\mathbf{K}}\check{\mathbf{K}}(\check{\mathbf{y}} - \check{\mathbf{H}}\hat{\mathbf{H}}\hat{\mathbf{x}}^b) \quad (4)$$

If we neglect the influence of the averaging kernel, this equation has the desired shape of Eq. (1), i.e. the sum of the prior value and of a linear function of model-minus-measurement misfits. However, to follow the optimal estimation framework, we still need to be able to develop the product of the gain matrices consistently with Eq. (2), i.e. like (neglecting errors in the observation operators):

$$\mathbf{K} = \hat{\mathbf{B}}\hat{\mathbf{H}}^T \check{\mathbf{H}}^T (\check{\mathbf{H}}\hat{\mathbf{H}}\hat{\mathbf{B}}\hat{\mathbf{H}}^T \check{\mathbf{H}}^T + \check{\mathbf{R}})^{-1} \quad (5)$$

In practice, we see that:

$$\hat{\mathbf{K}}\check{\mathbf{K}} = \hat{\mathbf{B}}\hat{\mathbf{H}}^T (\hat{\mathbf{H}}\hat{\mathbf{B}}\hat{\mathbf{H}}^T + \hat{\mathbf{R}}')^{-1} \check{\mathbf{B}}\check{\mathbf{H}}^T (\check{\mathbf{H}}\check{\mathbf{B}}\check{\mathbf{H}}^T + \check{\mathbf{R}})^{-1} \quad (6)$$

Equations (5)–(6) can be made consistent provided

$$\check{\mathbf{B}} = \hat{\mathbf{H}}\hat{\mathbf{B}}\hat{\mathbf{H}}^T \quad (7)$$

and (using Eq. 7)

$$(\hat{\mathbf{H}}\hat{\mathbf{B}}\hat{\mathbf{H}}^T + \hat{\mathbf{R}}')^{-1} \hat{\mathbf{H}}\hat{\mathbf{B}}\hat{\mathbf{H}}^T = \mathbf{I} \quad (8)$$

Equation (7) simply expresses consistency between the prior error statistics: the uncertainty of the retrieval prior should correspond to the uncertainty of the flux prior projected in the profile space. This condition is not achieved by current satellite retrieval algorithms because they artificially maximize the measurement contribution in

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the retrievals through the use of very large prior error variances (see Sect. 2.1). However, if enough intermediate variables were saved by the retrieval schemes, it would be possible to reconstruct the retrievals with a different prior error covariance matrix $\check{\mathbf{B}}$.

Equation (8) can obviously only be satisfied if the retrieval error variances are negligible compared to the flux prior error variances projected in the same space (which would actually relax the previous requirement as well). Typically, the standard deviation of the uncertainty (1σ) in the de-trended columns simulated by free models is not larger than a couple of ppm, at least for broad scale statistics (Chevallier and O'Dell, 2013; Peng et al., 2015), i.e. about the current GOSAT retrieval uncertainty (Oshchepkov et al., 2013). Note that the situation is more favourable when considering TCCON retrievals, because of their better precision.

As a consequence, the effective gain matrix $\hat{\mathbf{K}}\check{\mathbf{K}}$ significantly differs from the optimal one for GOSAT, resulting in a wrong balance between prior flux information and measured radiances. Overall, $\check{\mathbf{K}}$ pulls too much towards the measured radiances and $\hat{\mathbf{K}}$ pulls too much towards the prior. This suboptimality very likely flaws the 4-D information flow from the radiance measurements to the surface flux estimates. Further, the sub-optimality of $\check{\mathbf{K}}$ also affects the retrieval averaging kernel that is part of $\hat{\mathbf{H}}$, meaning that the model-data-misfits in Eq. (4) are not computed correctly, for instance because the retrieval averaging kernel would not peak low enough in the vertical.

The situation complicates even further if we account for the facts that inversion systems assimilate the retrieved profiles as vertical integrals (because $X\text{CO}_2$ is less sensitive to vertical transport model errors than the CO_2 profile), that these vertical integrals are empirically bias-corrected (thereby implicitly re-introducing $\check{\mathbf{x}}^b$ that had been neutralised by the use of averaging kernels, in the equations) and that $\check{\mathbf{H}}$ and $\hat{\mathbf{H}}$ are imperfect operators, the uncertainty of which should be reported in $\check{\mathbf{R}}$, following Eq. (5). The need to report all observation operator uncertainties in $\check{\mathbf{R}}$ means that retrieval configuration should in principle be tailored to the retrieval end-application, i.e. to the precision of the observation operator that is used in this end-application. For flux inversion, the transport model uncertainty in $X\text{CO}_2$ space is about 0.5 ppm (1σ ,

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diances and to generate a sizeable increment accordingly. By comparison, the model variability for a given increment size over the four years ranges between 3 and 4 ppm (1σ), the prior variability is about 3 ppm and the retrieval variability ranges between 3 and 7 ppm. The standard deviation that uses XCO_2^a is 1.1 ppm for small increments. It smoothly increases to 4 ppm for retrieval increments of size 6 ppm: it is systematically larger than the standard deviation that uses XCO_2^b (despite a smaller mean difference). The standard deviation that uses $XCO_2^{a,c}$ is also 1.1 ppm for small increments and is also systematically larger than the standard deviation that uses XCO_2^b , but it performs better than XCO_2^a . The worse standard deviation of the misfit of XCO_2^a and $XCO_2^{a,c}$ to the model compared to XCO_2^b cannot be explained by a common lack of variability in the model and in XCO_2^b , because thinning the retrievals (for instance by keeping only one retrieval every nine model grid boxes for a given day) only marginally changes the figure (not shown).

The fact that the standard deviation smoothly increases with increment size suggests that the increment size is systematically overestimated. Figure 6 presents an empirical test where we halve the retrieval increments, without any bias correction: we call $XCO_2^{a,r} = XCO_2^b + (XCO_2^a - XCO_2^b)/2$ the result. The reduction is seen to cancel most of the dependency of the statistics of the observation-minus-model misfits to the increment size: the standard deviation and the mean are then stable around 1.1 and -0.3 ppm, respectively for increments up to 4 ppm without any bias correction. The standard deviation is systematically better than for XCO_2^b , which shows added value brought by the radiance measurements, in contrast to the previous results.

For the medium-gain retrievals (Fig. 7) and for the ocean glint retrievals (Fig. 8), the standard deviation of the misfits using $XCO_2^{a,c}$ is not significantly larger than that using XCO_2^b .

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Table 2. Continued.

Locality (identifier)	Period	Source
Pic du Midi, FR (PDM)	2001–2013	LSCE
Pacific Ocean, 0N (POC000)	1987–2011	NOAA/ESRL
Pacific Ocean, 5N (POCN05)	1987–2011	NOAA/ESRL
Pacific Ocean, 10N (POCN10)	1987–2011	NOAA/ESRL
Pacific Ocean, 15N (POCN15)	1987–2011	NOAA/ESRL
Pacific Ocean, 20N (POCN20)	1987–2011	NOAA/ESRL
Pacific Ocean, 25N (POCN25)	1987–2011	NOAA/ESRL
Pacific Ocean, 30N (POCN30)	1987–2011	NOAA/ESRL
Pacific Ocean, 5S (POCS05)	1987–2011	NOAA/ESRL
Pacific Ocean, 10S (POCS10)	1987–2011	NOAA/ESRL
Pacific Ocean, 15S (POCS15)	1987–2011	NOAA/ESRL
Pacific Ocean, 20S (POCS20)	1987–2011	NOAA/ESRL
Pacific Ocean, 25S (POCS25)	1987–2011	NOAA/ESRL
Pacific Ocean, 30S (POCS30)	1987–2011	NOAA/ESRL
Pacific Ocean, 35S (POCS35)	1987–2011	NOAA/ESRL
Palmer Station, Antarctica, US (PSA)	1979–2013	NOAA/ESRL
Point Arena, California, US (PTA)	1999–2011	NOAA/ESRL
Puy de Dome, FR (PUY)	2001–2013	LSCE
Ragged Point, BB (RPB)	1987–2013	NOAA/ESRL
South China Sea, 3N (SCSN03)	1991–1998	NOAA/ESRL
South China Sea, 6N (SCSN06)	1991–1998	NOAA/ESRL
South China Sea, 9N (SCSN09)	1991–1998	NOAA/ESRL
South China Sea, 12N (SCSN12)	1991–1998	NOAA/ESRL
South China Sea, 15N (SCSN15)	1991–1998	NOAA/ESRL
South China Sea, 18N (SCSN18)	1991–1998	NOAA/ESRL
South China Sea, 21N (SCSN21)	1991–1998	NOAA/ESRL
Mahe Island, SC (SEY)	1980–2013	NOAA/ESRL
Southern Great Plains, Oklahoma, US (SGP)	2002–2013	NOAA/ESRL
Shemya Island, Alaska, US (SHM)	1985–2013	NOAA/ESRL
Ship between Ishigaki Island and Hateruma Island, JP (SIH)	1993–2005	WDCGG/Tohoku University
Shetland, Scotland, GB (SIS)	1992–2003	WDCGG/CSIRO
Tutuila, American Samoa (SMO)	1979–2013	NOAA/ESRL
South Pole, Antarctica, US (SPO)	1979–2013	NOAA/ESRL
Ocean Station M, NO (STM)	1980–2009	NOAA/ESRL
Summit, GL (SUM)	1997–2013	NOAA/ESRL
Syowa Station, Antarctica, JP (SYO)	1986–2013	NOAA/ESRL
Tae-ahn Peninsula, KR (TAP)	1991–2013	NOAA/ESRL
Tierra Del Fuego, Ushuaia, AR (TDF)	1994–2013	NOAA/ESRL
Trinidad Head, California, US (THD)	2002–2013	NOAA/ESRL
Tromelin Island, F (TRM)	1998–2007	LSCE
Wendover, Utah, US (UTA)	1993–2013	NOAA/ESRL
Ulaan Uul, MN (UUM)	1992–2013	NOAA/ESRL
Sede Boker, Negev Desert, IL (WIS)	1995–2013	NOAA/ESRL
Sable Island, CA (WSA)	1979–2012	WDCGG/EC
Mt. Waliguan, CN (WLG)	1990–2013	NOAA/ESRL
Western Pacific Cruise (WPC)	2004–2013	NOAA/ESRL
Ny-Alesund, Svalbard, Norway and Sweden (ZEP)	1994–2013	NOAA/ESRL

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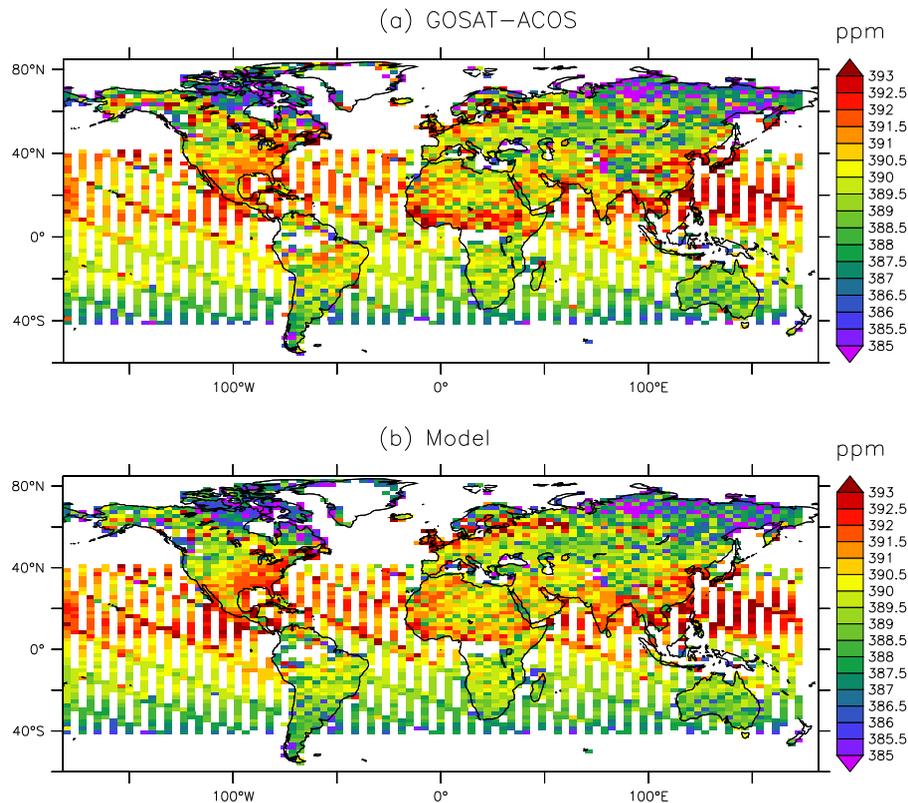
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Figure 1. (a) Mean ACOS-GOSAT bias-corrected retrievals in the model grid over 4 years (June 2009–May 2013). (b) Corresponding mean CO₂ 4-D field associated to the MACC CO₂ inversion (computed using the averaging kernels and the prior profiles of the retrievals).

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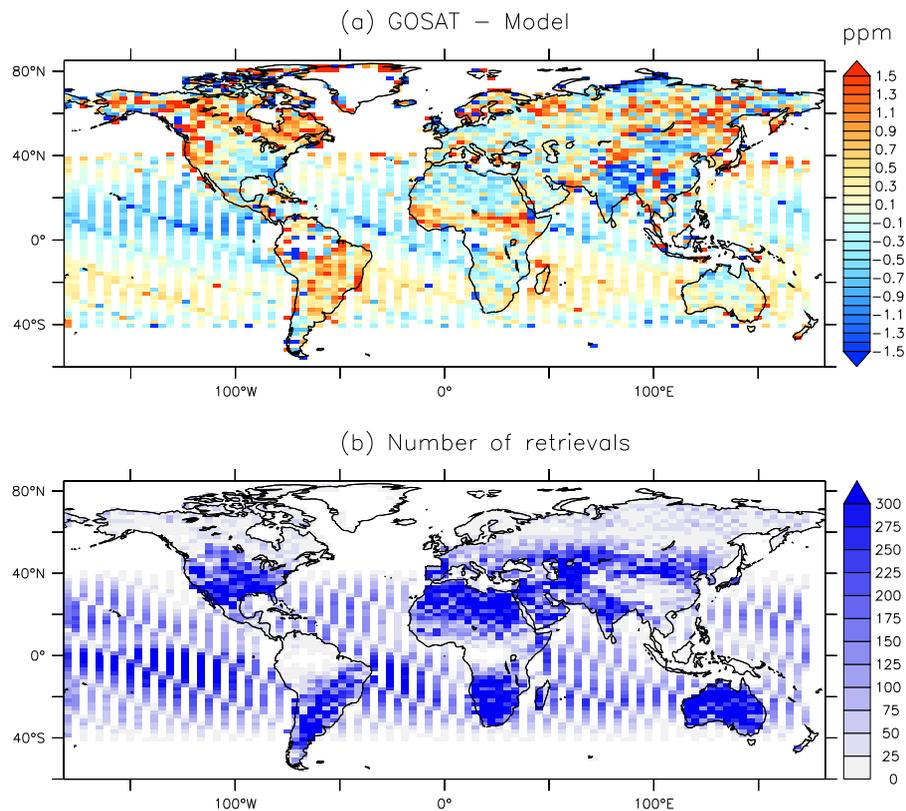


Figure 2. (a) Mean difference between the maps of Fig. 1 (retrievals minus model). (b) Corresponding number of retrievals.

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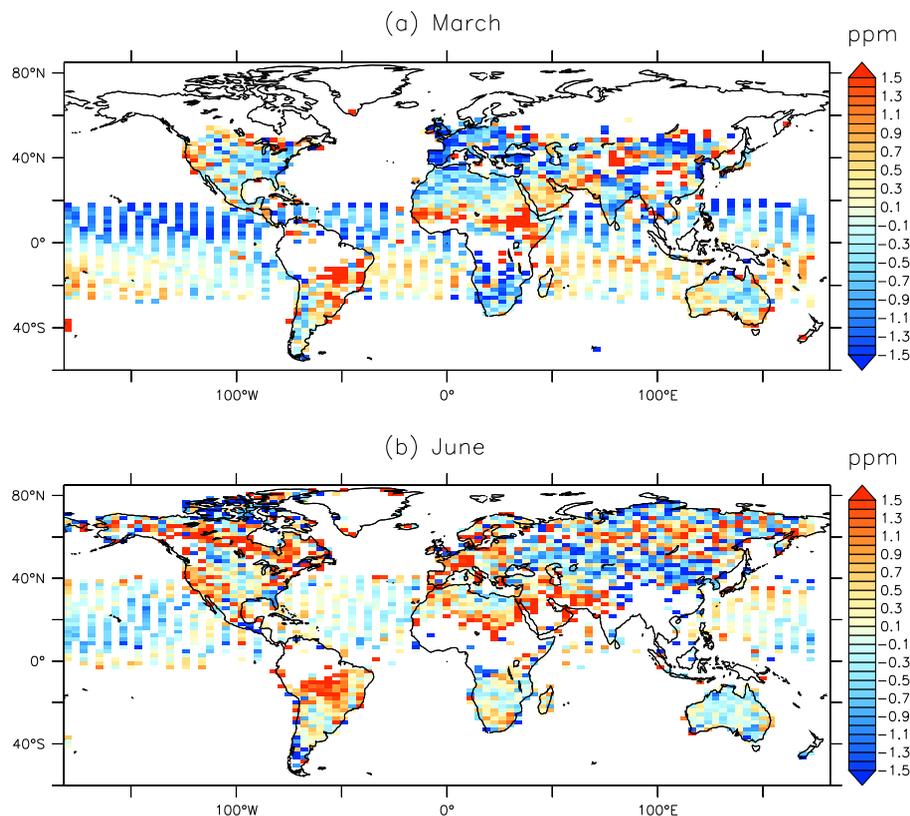
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Figure 3. Same as Fig. 2a (retrievals minus model), but focussing on the months of March and June.

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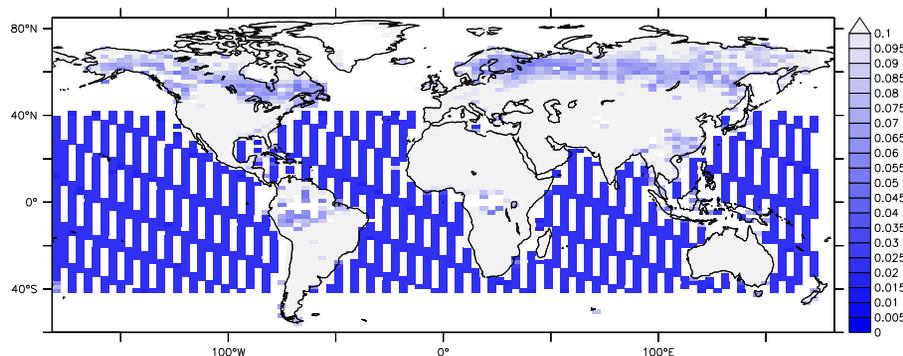


Figure 4. Mean surface albedo retrieved in the strong CO₂ band by ACOS-GOSAT in the model grid over 4 years (June 2009–May 2013). The blue scale focuses on the values less than 0.1.

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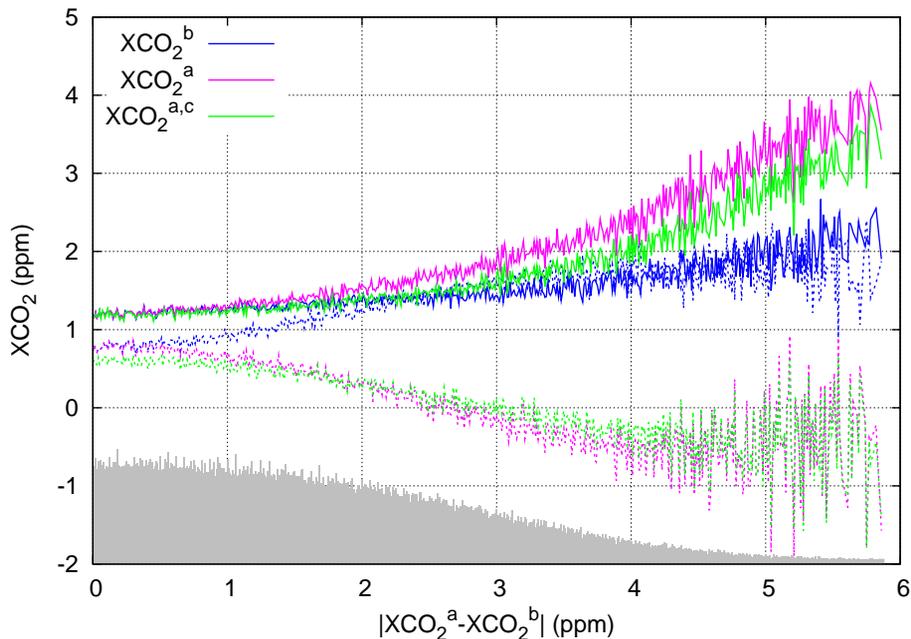


Figure 5. Mean and standard deviation of the retrieval-minus-model misfits between June 2009 and May 2013 for the high-gain mode retrievals over land as a function of the retrieval increment size. The statistics are also shown for the prior-minus-model misfit. The model values are raw pressure-weighted columns and do not account for the averaging kernels in order not to correlate the two axes (in practice, using the averaging kernels actually does not significantly affect the standard deviations shown). The grey shade shows the distribution of the retrieval density (axis not shown).

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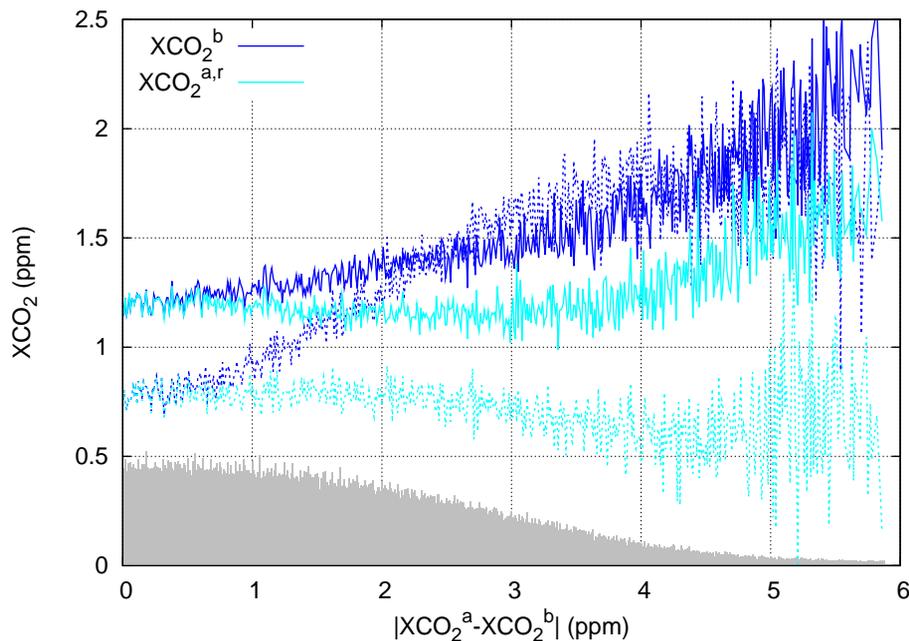


Figure 6. Same as Fig. 5 (high-gain mode over the lands) but we reduce the retrieval increment size by 50 % without any bias correction (we call $XCO_2^{a,r}$ the result). The abscissa shows the unperturbed increment.

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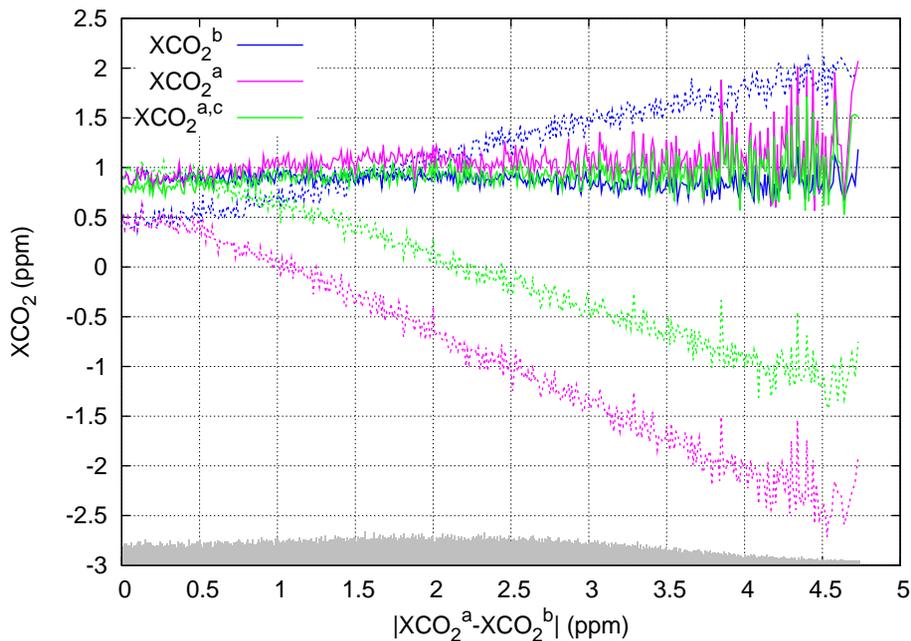
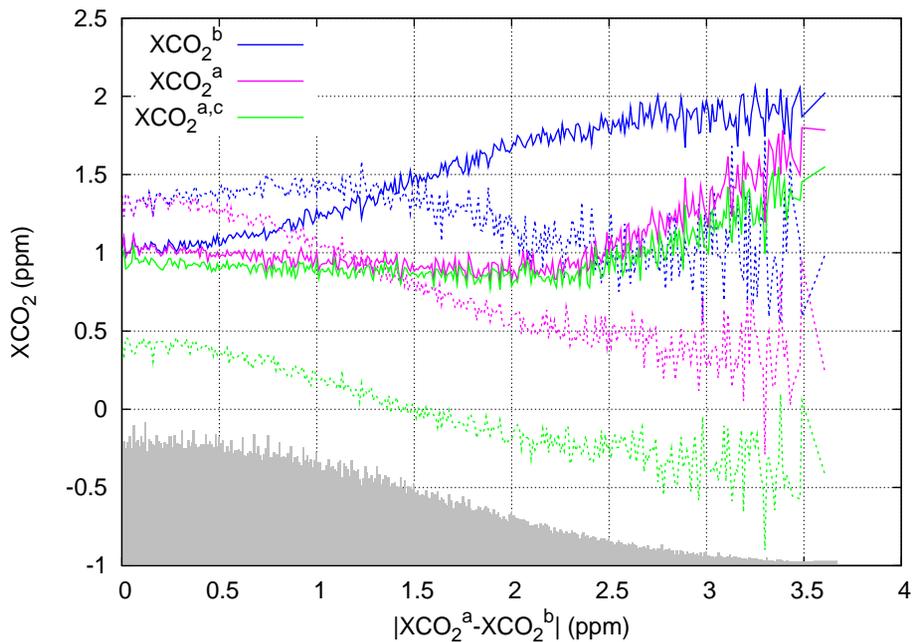
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Figure 7. Same as Fig. 5 for the medium-gain mode.

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**Figure 8.** Same as Fig. 5 for the glint mode over the ocean.[Title Page](#)[Abstract](#)[Introduction](#)[Conclusions](#)[References](#)[Tables](#)[Figures](#)[◀](#)[▶](#)[◀](#)[▶](#)[Back](#)[Close](#)[Full Screen / Esc](#)[Printer-friendly Version](#)[Interactive Discussion](#)